

# Models for Spatially Dependent Areal Data

## Models for Socio-Environmental Data

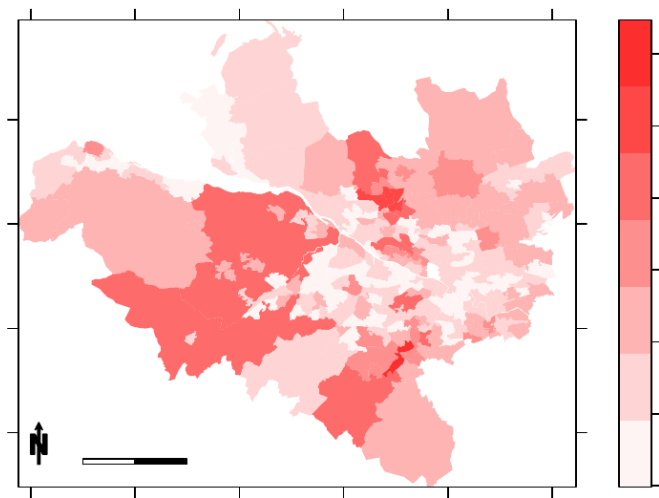
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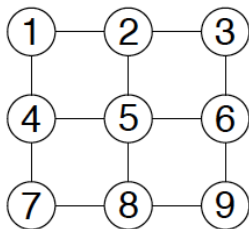


# Most ecological data are spatial

## Areal spatial processes



# Areal data and proximity



$$W = \begin{pmatrix} 0 & 1 & 0 & 1 & 0 & 0 & 0 & 0 & 0 \\ 1 & 0 & 1 & 0 & 1 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 & 1 & 0 & 0 & 0 \\ 1 & 0 & 0 & 0 & 1 & 0 & 1 & 0 & 0 \\ 0 & 1 & 0 & 1 & 0 & 1 & 0 & 1 & 0 \\ 0 & 0 & 1 & 0 & 1 & 0 & 0 & 0 & 1 \\ 0 & 0 & 0 & 1 & 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 1 & 0 & 1 & 0 & 1 \\ 0 & 0 & 0 & 0 & 0 & 1 & 0 & 1 & 0 \end{pmatrix}$$

Possibilities include, but are not limited to:

- ▶  $w_{ij} = 1$  if  $i, j$  share a common boundary (possibly a common vertex)
- ▶  $w_{ij} = 1$  for  $m$  nearest neighbors.

## Measures of regularity, clustering

Moran's I: similar to covariogram ( see spdep package ).

$$\hat{\mathbf{u}} = \mathbf{y} - \mathbf{X}\boldsymbol{\beta} \quad (1)$$

$$\mathbf{u} = \rho \mathbf{W}\mathbf{u} + \boldsymbol{\varepsilon}, \boldsymbol{\varepsilon} \text{ multivariate normal}(\mathbf{0}, \sigma_{\varepsilon}^2 \mathbf{I}) \quad (2)$$

Moran's I: (3)

$$I = \frac{\hat{\mathbf{u}}'}{||\hat{\mathbf{u}}||^2} \quad (4)$$

- ▶  $E(I) = -\frac{1}{n-1}$
- ▶  $I > E(I)$  implies clustering.
- ▶  $I < E(I)$  implies regularity.

## Modeling areal data

Two general types of *spatial autoregressive models*:

- ▶ Simultaneous autoregressive models (SAR): less common in Bayesian analysis. Some references at end of lecture.
- ▶ Conditional autoregressive models (CAR): The probability of values estimated at any given location are conditional on neighboring values.

# Conditional autoregressive model

$$\mathbf{y} = g(\boldsymbol{\theta}, \mathbf{X}) + \boldsymbol{\eta} + \boldsymbol{\varepsilon}$$

$$\boldsymbol{\eta} \sim \text{multivariate normal}(\mathbf{0}, \boldsymbol{\Sigma})$$

$$\boldsymbol{\Sigma} = \underbrace{\sigma^2(\mathbf{I} - \rho\mathbf{W})^{-1}\mathbf{I}}_{\text{covariance matrix}}$$

$$\boldsymbol{\varepsilon} \sim \text{multivariate normal}(\mathbf{0}, \sigma_\varepsilon^2\mathbf{I})$$

1.  $\rho$  is an autocorrelation parameter.
2. Proximity matrix  $\mathbf{W}$  must be symmetric.

## Alternative notation

$$\mathbf{y} \sim \text{multivariate normal}(g(\boldsymbol{\theta}, \mathbf{X}) + \boldsymbol{\eta}, \sigma_{\varepsilon}^2 \mathbf{I})$$

$$\boldsymbol{\eta} \sim \text{multivariate normal}(0, \boldsymbol{\Sigma})$$



# Conditional autoregressive model with row standardization

$$\mathbf{y} \sim \text{multivariate normal}(g(\boldsymbol{\theta}, \mathbf{X}) + \boldsymbol{\eta}, \sigma_{\varepsilon}^2 \mathbf{I})$$

$$\boldsymbol{\eta} \sim \text{multivariate normal}(0, \boldsymbol{\Sigma})$$

$$\boldsymbol{\Sigma} = (\text{diag}(\mathbf{W}\mathbf{1}) - \rho \mathbf{W})^{-1}$$

- ▶ Row standardization assures that  $|\rho| < 1$
- ▶  $\mathbf{1}$  is a column vector of 1's.
- ▶  $\mathbf{W}\mathbf{1}$  is the sums of the rows
- ▶  $\text{diag}(\mathbf{W}\mathbf{1})$  is a matrix with the sums of the rows on the diagonal and zeros elsewhere.
- ▶ Equivalent to dividing each element of  $\mathbf{W}$  by the sum of the rows to obtain  $\mathbf{W}_+$  and using  $\boldsymbol{\Sigma} = \sigma^2(\mathbf{I} - \rho \mathbf{W}_+)^{-1}$
- ▶  $\rho \sim \text{Beta}(18, 2)$  to favor values close to 1
- ▶  $\sigma^2 \sim \text{IG}(r, q)$ .

# CAR for non-negative observations

Let  $\Sigma = \sigma^2 \text{diag}(\mathbf{W}\mathbf{1} - \rho\mathbf{W})^{-1}$

$$g(\boldsymbol{\beta}, \mathbf{X}) = \exp(\mathbf{X}\boldsymbol{\beta}) \quad (5)$$

$$\log(y) \sim \text{multivariate normal}(\log(g(\boldsymbol{\beta}, \mathbf{X})) + \boldsymbol{\eta}, \sigma_\varepsilon^2 \mathbf{I}) \quad (6)$$

## CAR for counts

Let  $\Sigma = \sigma^2 \text{diag}(\mathbf{W}\mathbf{1} - \rho\mathbf{W})^{-1}$

$$g(\boldsymbol{\beta}, \mathbf{X}) = \exp(\mathbf{X}\boldsymbol{\beta}) \quad (7)$$

$$y_i \sim \text{Poisson}(\lambda_i) \quad (8)$$

$$\log(\boldsymbol{\lambda}) \sim \text{multivariate normal}(\log(g(\boldsymbol{\beta}, \mathbf{X})) + \boldsymbol{\eta}, \boldsymbol{\sigma}_\varepsilon^2 \mathbf{I}) \quad (9)$$

## CAR for binary observations

Let  $\Sigma = \sigma^2 \text{diag}(\mathbf{W}\mathbf{1} - \rho\mathbf{W})^{-1}$

$$g(\boldsymbol{\beta}, \mathbf{X}) = \text{logit}^{-1}(\mathbf{X}\boldsymbol{\beta}) \quad (10)$$

$$y_i \sim \text{Bernoulli}(p_i) \quad (11)$$

$$\text{logit}(\mathbf{p}) \sim \text{multivariate normal}(\text{logit}(g(\boldsymbol{\beta}, \mathbf{X})) + \boldsymbol{\eta}, \sigma_\epsilon^2 \mathbf{I}) \quad (12)$$

## Take home

- ▶ Data taken over time or space are likely to be structured by physical and biological processes.
- ▶ Our deterministic model may account for this structure. However, if the *residuals* show correlation over time and/ or space, then we are obliged to model their covariance to assure that iid assumptions are met.
- ▶ Doing so requires estimating only a few more parameters, in most cases one or two, relative to the aspatial model.
- ▶ Fitting spatial models is computationally challenging.
- ▶ Deciding whether to use a spatial or aspatial model should probably be treated as a problem in model selection.

## Further study

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- ▶ Finley, A. O., S. Banerjee, and B. P. Carlin. 2007. spBayes: An R package for univariate and multivariate hierarchical point-referenced spatial models. *Journal of Statistical Software* 19.
- ▶ Finley, A. O., S. Banerjee, and A. E. Gelfand. 2015. spBayes for Large Univariate and Multivariate Point-Referenced Spatio-Temporal Data Models. *Journal of Statistical Software* 63:1-28.
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- ▶ Cressie, N., and C. K. Wikle. 2011. *Statistics for spatio-temporal data*. Wiley.